#### Survival Estimation: Known-Fate Models

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#### Some Known-fate Models

- Binomial survival model
- Nest success
- Radiotelemetry data
- Study design

## Binomial Survival Model

• Follow *n* subjects, *x* of them survive with probability *s* 

$$f(x/n, s) = \binom{n}{x} s^{x} (1 - s)^{n - x}$$

$$\hat{s} = \frac{x}{n}; \hat{\text{var}}(\hat{s}) = \frac{\hat{s}(1 - \hat{s})}{n}$$

- Independent
- All detected
- No censoring

#### Mule Deer Example

|           | Number<br>Released | Alive | Dead | Other |
|-----------|--------------------|-------|------|-------|
| Treatment | 61                 | 19    | 38   | 4     |
| Control   | 59                 | 21    | 38   | 0     |

Treatment group:

$$\hat{s} = \frac{19}{57} = 333; \hat{var}(\hat{s}) = \frac{0.333(1 - 0.333)}{57} = 0.003899$$

#### **Nest Success**

- Important component of reproductive rate for may species (e.g., birds, many reptiles)
- Definition:
  - Pr (new nest succeeds to produce at least one hatchling or fledgling)
- Used with mean hatchlings/fledglings per successful nest

### Inference About Nest Success: Example

- Random sample of n bird nests
  - y = 1 represents nest success, y = 0 represents nest failure
  - 10 nests are observed daily until success or failure
  - outcome  $\underline{y} = \{1, 0, 1, 1, 1, 0, 0, 0, 1, 1\}$
- Assume: nest fates are independent and identically distributed with unknown probability p of nest success.
   We can model the probability associated with success
- Pr(x successes out of 10 nesting attempts) is binomial:

$$f(x/p) = {10 \choose x} p^{x} (1-p)^{10-x} \qquad L(p/x=6) = {10 \choose 6} p^{6} (1-p)^{4}$$
$$\frac{6}{p} \cdot \frac{4}{1-p} = 0, \qquad \hat{p} = 0.6. \quad \text{Is this OK?}$$

### Nest Studies and the Mayfield Method

- · Nest searches estimate nest success
  - Many nests encountered late in nesting phase
  - Positive bias in survival
    - · "Early" nests have more survival days
    - · Chance of failure related to # of days remaining until success
- Harold Mayfield (amateur birder) proposed idea
  - Estimate a daily nest survival probability, s, from sample
  - Use knowledge of number of days in entire nesting cycle, *d*, to estimate nesting success, *S*, as:

$$\hat{S} = \hat{S}^d$$

#### Study Design

- · Search for nests
  - Encounter active nest: revisit at intervals until
    - Failure
    - Fledging
- Intervals l = 1,...,L
- n<sub>l</sub>= Number of intervals of duration *l* for which fate was recorded
  - =  $n_{ls} + n_{lf}$  (n with success, n with failures)

## Mayfield's Estimator

 MLE for case where all nests visited each interval (every day)

$$\hat{s} = \frac{n_s}{n}$$
;  $\hat{var}(\hat{s}) = \frac{\hat{s}(1-\hat{s})}{n}$ 

• *n* = Number of intervals of duration 1 for which fate was recorded

=  $n_s + n_f$  ( $n_s$  with successes,  $n_f$  with failures)

### More General Approach

CA-A--A---A----A

CA-----A

CA----A----N

CA--A---A----A

CA = discovered nest; A = active; N = not active

 n<sub>l</sub> = Number of intervals of duration l for which fate was recorded

=  $n_{ls} + n_{lf}$  ( $n_{ls}$  with success,  $n_{lf}$  with failures)

### Example L=6

| Interval<br>length  | $n_{l.}$ | $n_{ls}$ |
|---------------------|----------|----------|
| 1 (s)               | 81       | 79       |
| $2 (s^2)$           | 30       | 29       |
| $3 (s^3)$           | 61       | 61       |
| 4 (s <sup>4</sup> ) | 37       | 36       |
| 5 (s <sup>5</sup> ) | 11       | 10       |
| 6 (s <sup>6</sup> ) | 21       | 19       |

 $\hat{\mathbf{s}} = 0.989, (0.981, 0.997)$ 

#### Estimation

• Binomial for each *l*, with survival constant but taken to appropriate power

$$f(n_{ls}/n_{l.},s) = \prod_{l=1}^{L} \frac{n_{l.}!}{n_{ls}! n_{lf}!} (s^{l})^{n_{ls}} (1-s^{l})^{n_{f}}$$

• MARK can be used to estimate s

#### Assumptions

- · Rates constant
  - Accommodate variation through stratification
- Visits recorded
- Pr(s) not influenced by observer
- Pr(visit) independent of Pr(survival)

# Untrastructural model for covariates

· Associate factors with success

$$s = \frac{e^{(\beta_0 + \sum_{j} \beta_j x_j)}}{1 + e^{(\beta_0 + \sum_{j} \beta_j x_j)}}$$

# Nest Success Models: Extensions/Advances

- Pollock and Cornelius (1988)
  - Use age of nest when first encountered to estimate age (stage) specific survival of nests
- Natarjan and McCulloch (1999)
  - Random-effects model (heterogeneity)
- Dinsmore et al. (2002)
  - Age-specific survival to approximate heterogeneity
- Rotella et al. (2004)
  - Review of approaches
- Etterson et al. (2007)
  - Partioning risk (causes of failure)

#### Design Issues: Nest Success

- Can predict n of samples (nests) needed
- · Tension between more nests and more visits
  - Fewer visits & more nests = increased precision
  - Fewer visits = less information on stage transitions and fledging
- · Need to worry about time of fledging
  - If miss final days, confuse mortality for success
  - Possible positive bias in survival rate
  - Age nests, and time visits for first possible day of fledging; do not include exposure days after that date
  - More formal approaches: Stanley (2000), Stanley and Grubb (2004)

#### Radiotelemetry Studies

- · Unlike nest studies, no natural endpoint
- · Heisey and Fuller (1985) approach
  - Assumes that survival may vary among intervals  $L_i$ , i=1,...,k (over time)
  - x<sub>i</sub> Number of "transmitter" days
  - d<sub>i</sub> Number of deaths in interval

$$\hat{\mathbf{S}}_i = \frac{(\mathbf{x}_i - \mathbf{d}_i)}{\mathbf{x}_i} \qquad \hat{\mathbf{S}} = \prod_{i=1}^k \hat{\mathbf{S}}_i^{L_i}$$

Censuring can occur (varies number of transmitter days)

#### Source-Specific Mortality

- $m_{ij}$  = Pr( animal alive during a day in interval i dies during the day from mortality source j)
- $d_{ij}$  = number of deaths in interval i resulting from mortality source j

$$\hat{\boldsymbol{m}}_{ij} = \frac{(d_{ij})}{x_i}$$

• Probability that animal dies as result of source *j* during interval *i* is

$$\hat{M}_{ij} = \hat{m}_{ij} + \hat{S}_{ij} \hat{m}_{ij} + \hat{S}_{ij} \hat{m}_{ij} + \dots + \hat{S}_{ij} \hat{m}_{ij} + \dots + \hat{S}_{ij} \hat{m}_{ij}$$

$$= [\hat{m}_{ij} / (1 - \hat{S}_{i})](1 - \hat{S}_{i})$$

### Kaplan-Meier Nonparametric Survival Model

• Discrete hazard function h<sub>i</sub>;

Pr(die on day j / alive just prior to j)

- $-d_j$  = # of animals that die at time j, for observed times at death,  $t_1, t_2, \dots t_j$
- $-r_i = \#$  of animals at risk in just prior to  $t_i$

$$\hat{h_j} = \frac{d_j}{r_i}$$

• Empirical estimate of S(t);

Pr(survive until at least t)  $\hat{S}(t) = \prod_{i=1}^{t} (1 - \hat{h}_i)$ 

· Product limit estimator

#### Kaplan-Meier Nonparametric Survival Model

- Pollock et al (1989) noted that model and estimator can accommodate
  - Addition of animals
  - Temporary censoring
  - Permanent censoring
  - modify  $r_j$  to subtract deaths and censoring and add new animals

# Parametric and Nonparametric Survival Models

- Failure time analysis: huge field
  - Models of time until failure
  - Model instantaneous hazard rate h(t)
- · Earlier models:
  - Fixed time periods, binomial models
- Parametric models (model survival function)
  - Proportional hazards, Weibull
- · Kaplan-Meier Nonparametric Survival Model

#### Design Considerations I

- · Capture n animals
  - How many? Use binomial model for sample allocation
  - "Staggered entry" is OK
- · Postrelease adjustment period
  - Minimize possible behavioral and radio effects

#### Design Considerations II

- Periodically survey area, record fates (alive or dead), cause of death
  - Study area must be small enough to permit frequent surveys
  - Try to prevent censoring
  - Animals not encountered should be censored, and if later resighted should be considered as a new staggered entry\*
  - Censoring must be random and independent of fate
    - · Frequently violated in studies